Session1: Bayesian Decision theory & Mixture Model

HMM for TS classif & filtering

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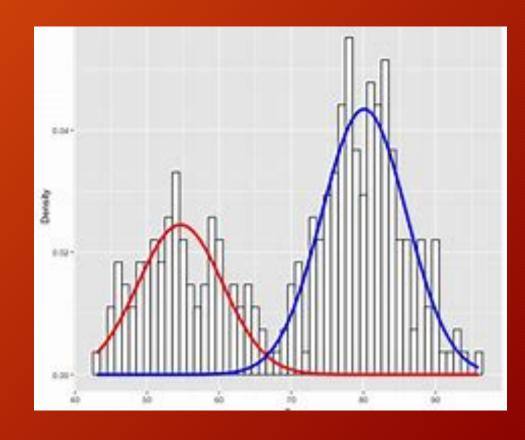
Outline of session 1 (8h)

3. Mixture Model

- Definition, simulations.
- Automatic parameters learning: EM principle
- Lab: Gaussian mixture model and Image processing

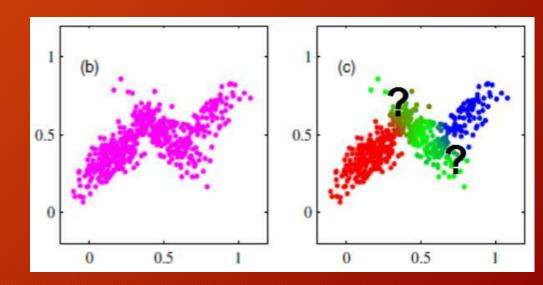
- 2D Gaussians
- 2D mixture and EM re-estimation update
- References and « Not seen! »
- 1h sitting-exam, date: October 1, 2019

- Definition, simulations
- Exemple of mixture: image processing
- Automatic parameters learning: the EM principle
- Exercise: Gaussian mixture model



In <u>statistics</u>, a <u>mixture model</u> is a <u>probabilistic model</u> for representing the presence of <u>subpopulations</u> within an overall population, without requiring that an observed data set should identify the sub-population to which an individual observation belongs.

Formally a mixture model corresponds to the <u>mixture</u> <u>distribution</u> that represents the <u>probability distribution</u> of observations in the overall population. However, while problems associated with "mixture distributions" relate to deriving the properties of the overall population from those of the subpopulations, "mixture models" are used to make <u>statistical</u> <u>inferences</u> about the properties of the sub-populations given only observations on the pooled population, without subpopulation identity information.



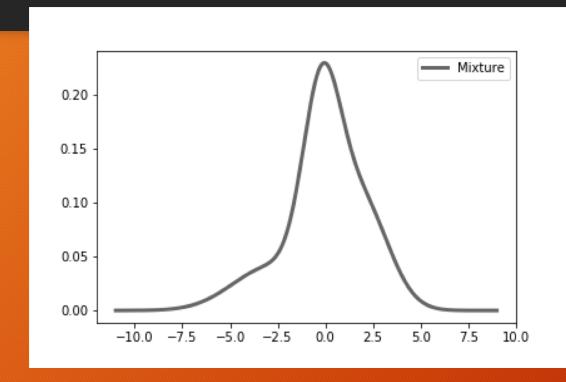
Suppose that we have a sample

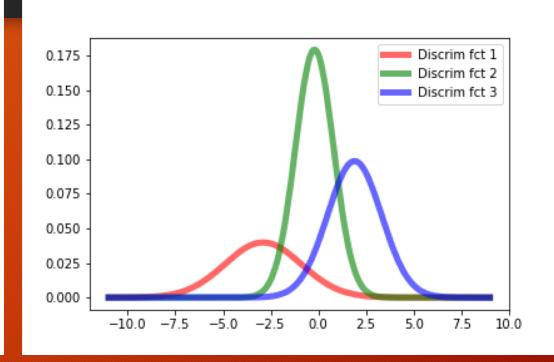
$$\mathbf{y} = \mathbf{y}_1^N = \{y_1, y_2, \dots, y_n, \dots, y_N\}$$

distributed according to a mixture of Gaussian distributions, so that all sample has the following with density:

$$P(Y_n = y_n) = f(y_n) = \sum_{k=1}^{K} \pi_k f_k(y_n)$$

A Gaussian mixture model is made with Gaussian f_k .





$$\mathcal{N}(\mu_1 = -2.9, \sigma_1 = 2)$$
 $\pi_1 = 0.10$

$$\mathcal{N}\left(\mu_2 = -0.2, \sigma_2 = 1\right)$$

$$\mathcal{N}\left(\mu_3 = 1.9, \sigma_3 = \sqrt{2}\right) \quad \pi_3 = 0.35$$

$$\pi_1 = 0.10$$

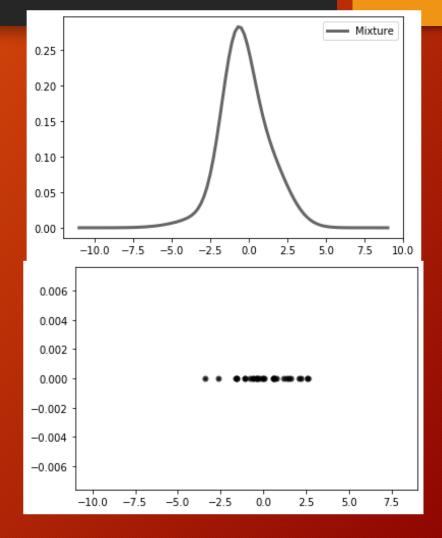
$$\pi_2 = 0.55$$

$$\pi_3 = 0.35$$

Question: How to draw a sample for a mixture?

$$P(Y_n = y_n) = f(y_n) = \sum_{k=1}^{K} \pi_k f_k(y_n)$$

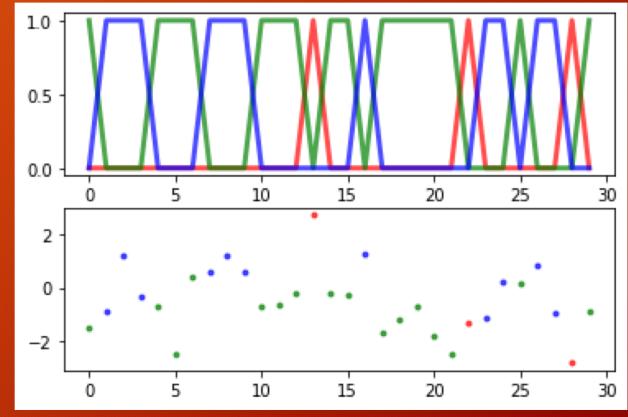
$$\mathcal{N}(\mu_1 = -2.9, \sigma_1 = 2)$$
 $\pi_1 = 0.10$ $\mathcal{N}(\mu_2 = -0.2, \sigma_2 = 1)$ $\pi_2 = 0.55$ $\mathcal{N}(\mu_3 = 1.9, \sigma_3 = \sqrt{2})$ $\pi_3 = 0.35$



Question: How to draw a sample for a mixture?

- 1. Sampling according to the a priori proba to get the class number.
- 2. Sampling according to the selected Gaussian.

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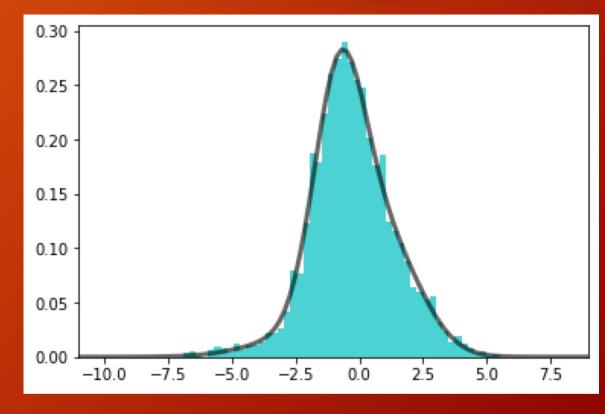
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For a sample 3000 data

In <u>statistics</u>, the <u>likelihood</u> expresses how probable a given set of <u>observations</u> is for different values of <u>statistical parameters</u>. It is equal to the <u>joint probability</u> <u>distribution</u> of the random sample evaluated at the given observations, and it is, thus, solely a function of parameters that index the <u>family</u> of those probability distributions.

For MM
$$\mathcal{L}_{\Theta}(oldsymbol{y}) = \prod_{n=1}^N \sum_{k=1}^K \pi_k \; f_k(y_n)$$



Sir Ronald Fisher

considered as a function of Θ .

Mapping from the <u>parameter space</u> to the <u>real line</u>, the likelihood function presents a peak, if it exists, which represents the combination of model parameter values that maximize the probability of drawing the sample actually obtained.

The procedure for obtaining these <u>arguments of the maximum</u> of the likelihood function is known as <u>maximum likelihood estimation</u>, which for computational convenience is usually done using the <u>natural logarithm</u> of the likelihood, known as the <u>log-likelihood function</u>.

Question: How to maximise $\mathcal{L}_{\Theta}(y)$?

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- Direct maximization is not possible.
- Solution: Expectation-Maximization (EM) algorithm

We define the 'joint likelihood'

$$\mathcal{H}_{\Theta}(\boldsymbol{y}, \boldsymbol{X}) = \prod_{n=1}^{N} \pi_{X_n} f_{X_n}(y_n) = \prod_{n=1}^{N} \sum_{k=1}^{K} \pi_k f_k(y_n) \mathbb{I}_{(X_n = k)}$$

This a random function of $oldsymbol{X} = oldsymbol{X}_1^N = \{X_1, X_2, \dots, X_n, \dots, X_N\}$

The EM algorithm: this is an iterative algorithm to estimate the maximum of the likelihood function, by computing iteratively two steps:

1. Expectation of the auxiliary function

where
$$\mathcal{Q}\left(\Theta;\Theta^{(\ell)}\right) = E\left[\ln \mathcal{H}_{\Theta}(\boldsymbol{y},\boldsymbol{X})|\Theta^{(\ell)}\right]$$

- Θ is the set of true parameters (we are looking for). $\Theta^{(\ell)}$ is the estimated parameters set at iteration ℓ .
- 2. Maximization of the auxiliary function

$$\Theta^{(\ell+1)} = \arg\max_{\Theta} \mathcal{Q}\left(\Theta; \Theta^{(\ell)}\right)$$

Properties of the EM algorithm (not proven)

1. Construction of a series of estimators for which the likelihood is increasing.

$$\mathcal{L}_{\Theta^{(\ell+1)}}(oldsymbol{y}) \geq \mathcal{L}_{\Theta^{(\ell)}}(oldsymbol{y})$$

The likelihood is always increasing (this is a sufficient condition to ensure the convergence of the EM algorithm).

Properties of the EM algorithm (not proven)

2. Convergence towards one of the (local) maxima of likelihood since we have

$$\left. \frac{\partial \mathcal{Q} \left(\Theta; \Theta^{(\ell)} \right)}{\partial \Theta} \right|_{\Theta = \Theta^{(\ell)}} = \left. \frac{\partial \mathcal{L}_{\Theta} (\boldsymbol{y})}{\partial \Theta} \right|_{\Theta = \Theta^{(\ell)}}$$

Initialization: Biernacki, C., Celeux, G. and Govaert, G. (2003). Choosing starting values for the EM algorithm for getting the highest likelihood in multivariate Gaussian mixture models. Computational Statistics and data analysis 41, 561-575.

The joint log-likelihood is written

$$\ln \mathcal{H}_{\Theta}(\boldsymbol{y}, \boldsymbol{X}) = \sum_{n=1}^{N} \sum_{k=1}^{K} \ln \left(\pi_k \ f_k(y_n) \right) \ \mathbb{I}_{(X_n = k)}$$

At iteration ℓ , the Gaussian pdfs write $f_k^{(\ell)}$ with parameters $\mu_k^{(\ell)}, \sigma_k^{(\ell)}$ and the auxiliary function writes

$$Q\left(\Theta;\Theta^{(\ell)}\right) = \sum_{n=1}^{N} \sum_{k=1}^{K} \ln\left(\pi_k^{(\ell)} f_k^{(\ell)}(y_n)\right) E\left[\mathbb{I}_{(X_n=k)}|\boldsymbol{y},\Theta^{(\ell)}\right]$$

$$Q\left(\Theta;\Theta^{(\ell)}\right) = \sum_{n=1}^{N} \sum_{k=1}^{K} \ln\left(\pi_k^{(\ell)} f_k^{(\ell)}(y_n)\right) E\left[\mathbb{I}_{(X_n=k)} | \boldsymbol{y}, \Theta^{(\ell)}\right]$$

with

$$E\left[\mathbb{I}_{(X_n=k)}|\boldsymbol{y},\boldsymbol{\Theta}^{(\ell)}\right] = p\left(X_n = k|\boldsymbol{y},\boldsymbol{\Theta}^{(\ell)}\right) = p\left(X_n = k|y_n,\boldsymbol{\Theta}^{(\ell)}\right)$$
$$= \gamma_n^{(\ell)}(k) = \frac{\pi_k^{(\ell)} f_k^{(\ell)}(y_n)}{\sum_{j=1}^K \pi_j^{(\ell)} f_j^{(\ell)}(y_n)}$$

<u>Gaussian mixture:</u> as an exercise, proof that EM-based re-estimation formulas for parameters of a MM can be written:

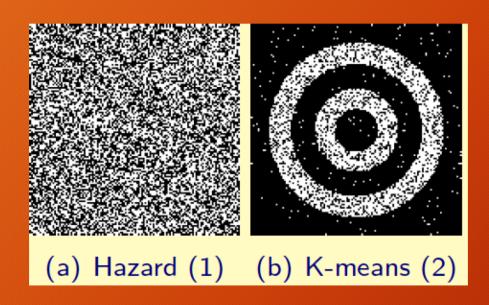
$$\pi_k^{(\ell+1)} = \frac{1}{N} \sum_{n=1}^N \gamma_n^{(\ell)}(k)$$

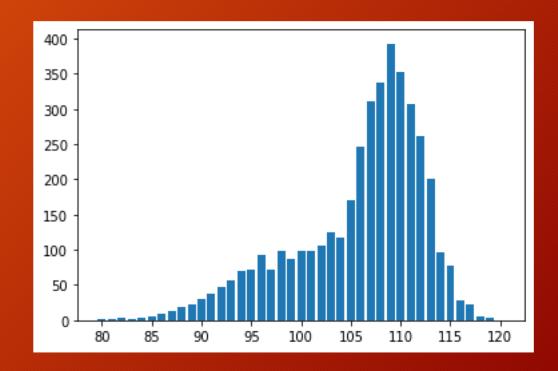
$$\mu_k^{(\ell+1)} = \frac{\sum_{n=1}^{N} \gamma_n^{(\ell)}(k) y_n}{\sum_{n=1}^{N} \gamma_n^{(\ell)}(k)}$$

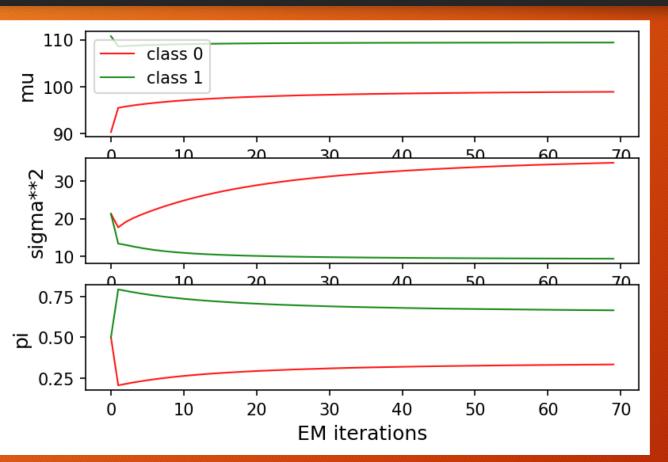
$$\sigma_k^{2,(\ell+1)} = \frac{\sum_{n=1}^{N} \gamma_n^{(\ell)}(k) \left(y_n - \mu_k^{(\ell+1)} \right)^2}{\sum_{n=1}^{N} \gamma_n^{(\ell)}(k)}$$

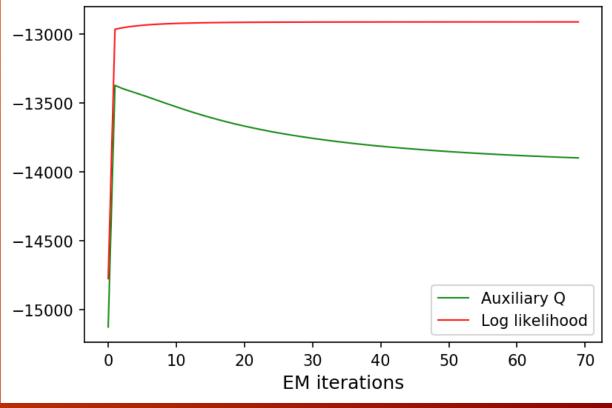
The EM algorithm looks for a local maxima of the likelihood: it requires the parameters to be initialized "not so far" from the true values.

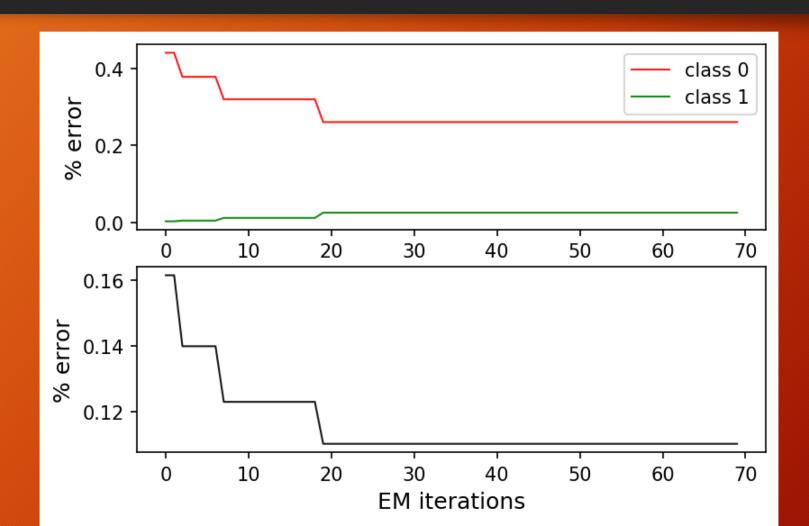
Any idea to initialize parameters?

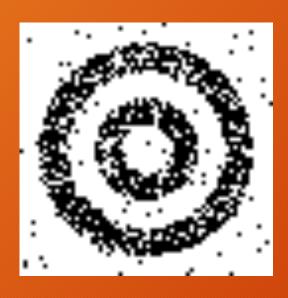






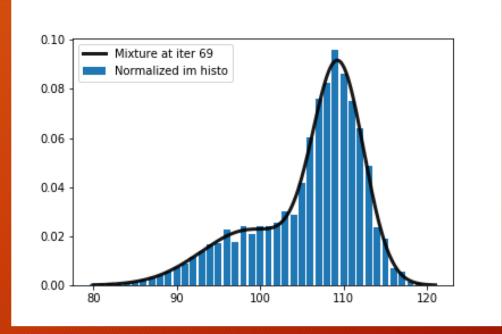






$$\xi_1 = 0.261$$

 $\xi_2 = 0.025$
 $\xi = 0.11$



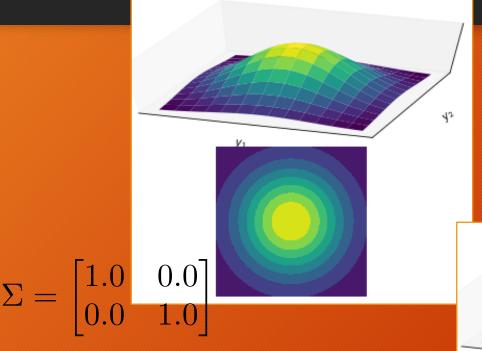
$$\pi_1 = 0.334$$
 $\mathcal{N}(\mu_1 = 98.85, \sigma_1 = 5.91)$
 $\pi_2 = 0.666$ $\mathcal{N}(\mu_2 = 109.39, \sigma_2 = 3.06)$

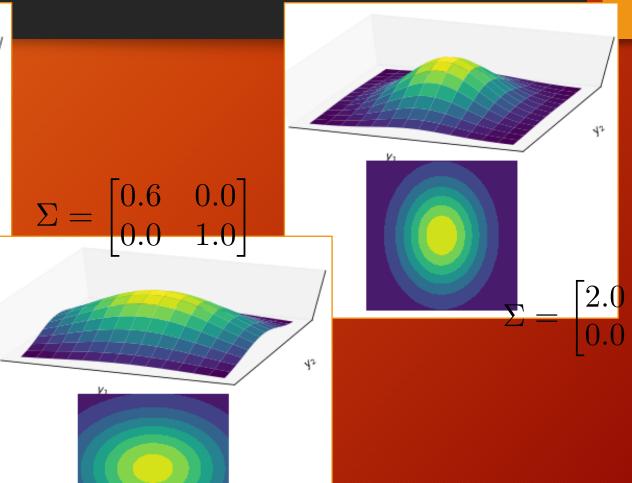
$$f(\boldsymbol{y}) = \sum_{k=1}^{K} \pi_k f_k(\boldsymbol{y}) \qquad \boldsymbol{y} = \begin{bmatrix} y_1 \\ y_2 \end{bmatrix}$$

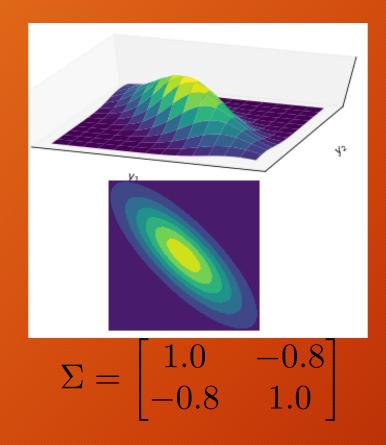
2D Gaussian

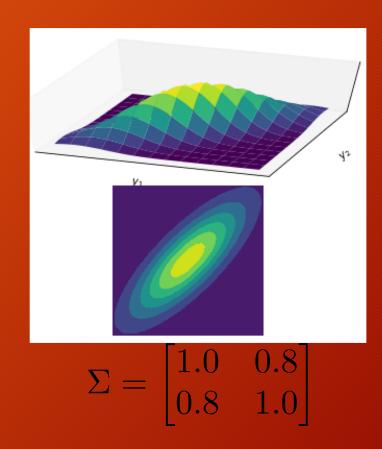
$$f(\boldsymbol{y}) = \frac{1}{\sqrt{(2\pi)^k |\Sigma|}} e^{-\frac{1}{2}(\boldsymbol{y} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\boldsymbol{y} - \boldsymbol{\mu})}$$

$$oldsymbol{\mu} = egin{bmatrix} \mu_1 \ \mu_2 \end{bmatrix} \quad oldsymbol{\Sigma} = egin{bmatrix} \sigma_1^2 &
ho\sigma_1\sigma_2 \
ho\sigma_1\sigma_2 & \sigma_2^2 \end{bmatrix}$$

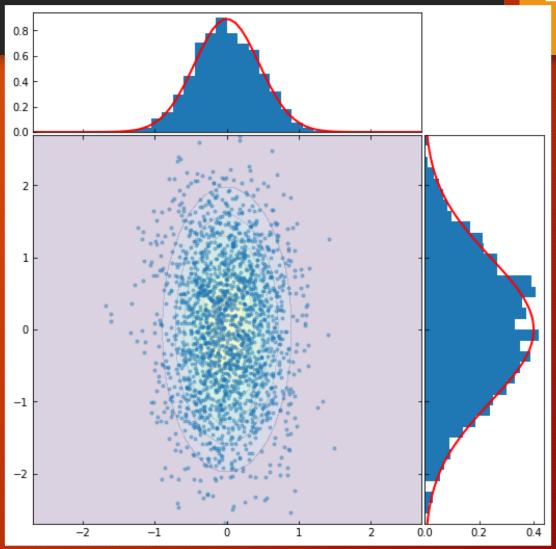








Margins, conditional laws, empirical estimation of paramters

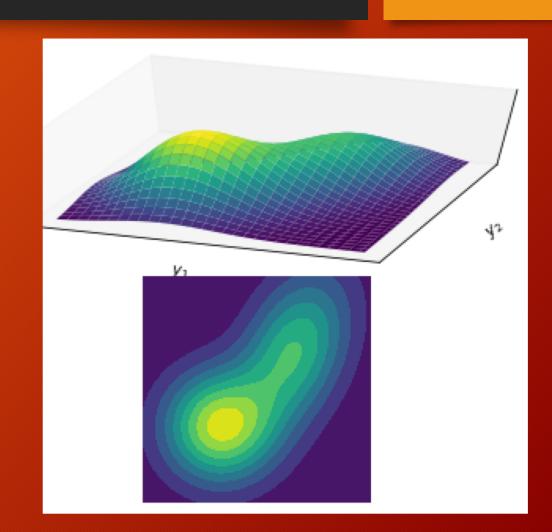


$$f(\mathbf{y}) = \sum_{k=1}^{\infty} \pi_k f_k(\mathbf{y})$$

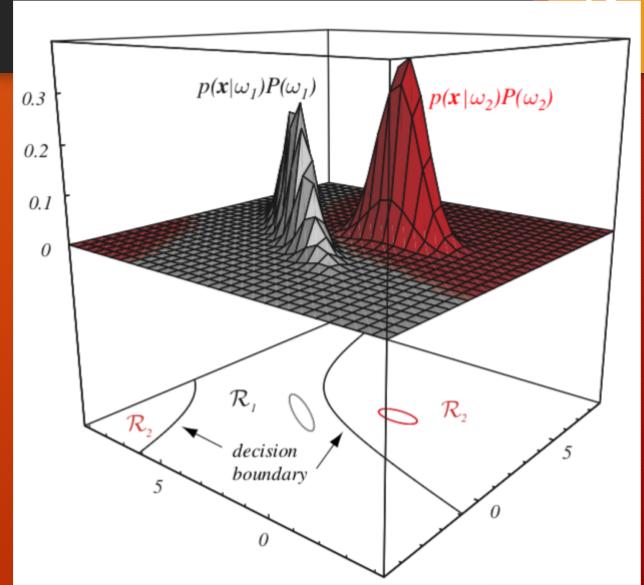
$$\pi_1 = \pi_2 = \frac{1}{2}$$

$$\boldsymbol{\mu}_1 = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \qquad \boldsymbol{\mu}_2 = \begin{bmatrix} 2 \\ 2 \end{bmatrix}$$

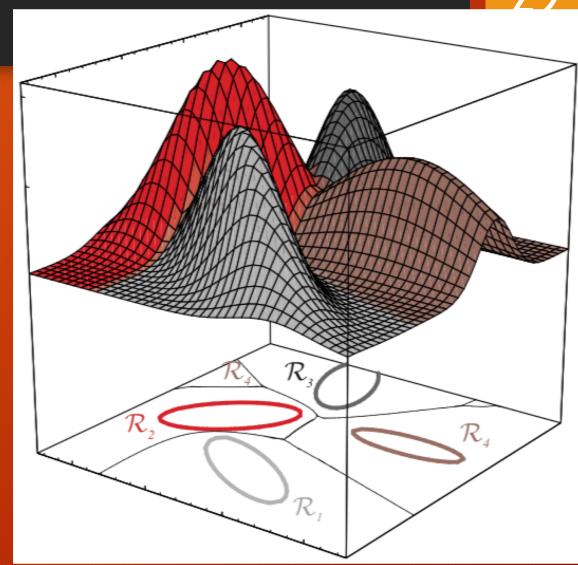
$$\boldsymbol{\Sigma}_1 = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \qquad \boldsymbol{\Sigma}_2 = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 2 \end{bmatrix}$$



Decision boundary can be complex!!! (i.e. not always linear)



Multi-class decision boudaries



EM-based re-estimation formulas for parameters of a 2D MM are essentially the same:

$$\pi_k^{(\ell+1)} = \frac{1}{N} \sum_{n=1}^N \gamma_n^{(\ell)}(k)$$

$$\boldsymbol{\mu}_k^{(\ell+1)} = \frac{\sum\limits_{n=1}^N \gamma_n^{(\ell)}(k) \; \boldsymbol{y}_n}{\sum\limits_{n=1}^N \gamma_n^{(\ell)}(k)}$$

$$\boldsymbol{\Sigma}_{k}^{(\ell+1)} = \frac{\sum\limits_{n=1}^{N} \gamma_{n}^{(\ell)}(k) \; \left(\boldsymbol{y}_{n} - \boldsymbol{\mu}_{k}^{(\ell+1)}\right)^{T} \left(\boldsymbol{y}_{n} - \boldsymbol{\mu}_{k}^{(\ell+1)}\right)}{\sum\limits_{n=1}^{N} \gamma_{n}^{(\ell)}(k)}$$

References

- Theory and Use of the EM Algorithm By Maya R. Gupta and Yihua Chen, Book pdf
- The EM algorithm and related statistical models By Michiko Watanabe and Kazunori Yamaguchi, <u>Book pdf</u>
- Pattern classification by Richard O. Duda, Peter E. Hart and David G. Stork, 2015, <u>Book pdf</u>, <u>Slides of the book</u>
- Finite Mixture model By Geoffrey McLachlan and D. Peel, <u>Book pdf</u>
- Python library for MM: Pymix, sklearn.mixture

Not seen!

- Variations about EM
 GEM, CEM, SEM -- On-line EM, by O. Cappé
- Mixture of non-gaussian type:
 - M. of generalized hyperbolic distribution
 - M. of skew-normal distribution, M. of t-distribution
- Choosing the number of clusters via model selection criteria
 - **BIC:** Bayesian Information Criterion,
 - **AIK:** Akaike Information Criterion,
 - ICL: Integrated completed likelihood criterion

Biernacki, C., G. Celeux, and G. Govaert (2000). "Assessing a mixture model for clustering with the integrated completed likelihood". IEEe Trans. PAMI, Vol 22(7), pp. 719-725.